

SCHOOL OF BUSINESS, ECONOMICS AND LAW

Department of Economics School of Business, Economics and Law Gothenburg University

# Volatility forecasting using the GARCH framework on the OMXS30 and MIB30 stock indices

Bachelor Thesis 15 credits Winter 2008 Supervisor: Jens Madsen Author: Peter Johansson

## **Executive Summary**

- **Title** Volatility forecasting using the GARCH framework on the OMXS30 and MIB30 stock indices
- **Subject** Bachelor Thesis in Financial Economics (15 credits) at the School of Business, Economics and Law, Gothenburg University
- **Author** Peter Johansson
- **Supervisor** Jens Madsen, Department of Economics, Gothenburg University
- **Key Words** Volatility forecasting, Random Walk, Moving Average, Exponentially Weighted Moving Average, GARCH, EGARCH, GJR-GARCH, APGARCH, volatility model valuation, regression, information criterion
- **Problem** There are many models on the market that claim to predict changes in financial assets as stocks on the Stockholm stock exchange (OMXS30) and the Milano stock exchange index (MIB30). Which of these models gives the best forecasts for further risk management purposes for the period  $31<sup>st</sup>$  of October 2003 to 30<sup>th</sup> of December 2008? Is the GARCH framework more successful in forecasting volatility than more simple models as the Random Walk, Moving Average or the Exponentially Weighted Moving Average?
- **Purpose** The purpose of this study is to find and investigate different volatility forecasting models and especially GARCH models that have been developed during the years. These models are then used to forecast the volatility on the OMXS30 and the MIB30 indices. The purpose of this study is also to find the model or models that are best suited for further risk management purposes by running the models through diagnostics checks.
- **Method** Daily prices together with the highest and lowest prices during one trading day for the OMXS30 and MIB30 indices were collected from Bloomberg for the period 31<sup>st</sup> of October 2003 to 30<sup>th</sup> of December 2008. These data were then processed in Microsoft Excel and Quantitative Micro Software EViews to find the most successful model in forecasting volatility for each of the two indices. The forecasting was performed on-step ahead for the period  $1<sup>st</sup>$  of July 2008 to 30 $^{\text{th}}$  of December 2008.
- **Results** This study has examined the forecasting ability of various models on the OMXS30 and MIB30 indices. The models were the Random Walk (RW), the Moving Average (MA), the Exponentially Weighted Moving Average (EWMA), ARCH, GARCH, EGARCH, GJR-GARCH and APGARCH. The results suggest that the best performing model was EGARCH(1,1) for both indices.

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# <span id="page-5-0"></span>**1 Introduction**

The paragraphs in the introduction section give a brief description of the background, problem definition, delimitations and purpose of this study. The disposition of this report is also presented.

## <span id="page-5-1"></span>**1.1 Background**

Trading in the world's financial markets has increased dramatically over the years. As larger and larger volumes are handled by investors, financial institutions and their traders, it has become more important to handle and evaluate risk. Evidence in the past has showed that lack of risk handling and control systems can have a devastating outcome. An example of lack of risk evaluation is told by the outcomes of the United Kingdom's oldest merchant bank, Barings Bank. Here one person traded with financial instruments for vast amount of money without risk handling. The result was that the Barings Bank went bankrupt and the total losses exceeded one billion dollars.

Financial instruments are today more exposed to volatility when global markets are connected to each other and tend to shift rapidly. Financial instruments are also becoming more complex and harder to understand. Therefore investors are paying more interest to the actual risk involved and not only the payoff (Simons, 2000).

Another perspective of risk in classic portfolio theory says that investors can eliminate assetspecific risk by diversifying their portfolio by holding many different assets. Although the risk can be diversified away in this manner, the market does not reward such portfolio handling. Instead an investor should hold a combination of a risk-free asset and the market portfolio. The conclusion here is that investors should not waste resources as investors don't care about firm-specific risk (Christoffersen, 2003). Many investors have however learnt that firm-specific risk shall have a high priority. Firm-specific risk involves bankruptcy costs, taxes paid which influence the value of the firms, capital structure and the cost of capital. All these risks influence firm value on the market and have a direct influence on the volatility of their stock value.

Many models have been developed over the years to forecast the volatility and thus to give a framework to manage risk. The RiskMetrics model, also called the exponential smoother, was one of the most famous models developed by J P Morgan. Another famous model is the Random Walk. These models are easy to use but have some shortcomings. The disadvantages with RiskMetrics or the Random Walk have been to drive the development of new models to give a better prediction of future volatility. A set of new models that have been developed are called GARCH, generalized autoregressive conditional heteroskedasticity. These models have showed enough flexibility to accommodate specific aspects of individual assets, while still being simple to use (Christoffersen, 2003).

# <span id="page-5-2"></span>**1.2 Purpose**

Firstly, the purpose of this study is to find and investigate different volatility forecasting models, especially GARCH models, which have been developed over the years. These models are then used to predict the volatility on the Stockholm stock exchange index (OMXS30) and the Milano stock exchange index (MIB30). Secondly, the aim is to run these models through

diagnostics checks to determine the model or models that are best suited for further risk management purposes.

# <span id="page-6-0"></span>**1.3 Problem definition**

There are many models on the market that claim to predict changes in financial assets as stocks on the OMXS30 and MIB30 indices. Which of these models gives the best predictions for further risk management purposes for the period 31<sup>st</sup> of October 2003 to 30<sup>th</sup> of December 2008? Is the GARCH framework more successful in predicting volatility than more simple models such as the Random Walk, Moving Average or the Exponentially Weighted Moving Average models?

## <span id="page-6-1"></span>**1.4 Delimitations**

The study presented in this report is limited to volatility forecasting on the Stockholm stock exchange index (OMXS30) and the Milano stock exchange index (MIB30). Models studied are the Random Walk (RW), the Moving Average (MA), the exponentially weighted Moving Average (EWMA), ARCH, GARCH, EGARCH, GJR-GARCH and APGARCH. The estimations of the models are performed under the assumption that the returns of the stock indices follow the normal distribution. No estimation is used under the assumption of a Student-t distribution, Skewed Student-t distribution, Gereralized Error Distribution or any other distribution. Furthermore, processing of the forecasted data as Value at Risk (VaR) or Expected Shortfall (ES) is not performed in this study.

The time period used for the data from the stock indices range from  $31<sup>st</sup>$  of October 2003 to 30<sup>th</sup> of December 2008.

### <span id="page-6-2"></span>**1.5 Disposition**

The report starts with chapter [1](#page-5-0) introducing the subject, the purpose of the study, the problem definitions and the delimitations. It then continues with chapter [2](#page-6-3) explaining the methodology and approach to the study. In chapter [3](#page-7-1) the theoretical framework goes through the theories used to retrieve the results. Chapter [4](#page-19-2) describes the data and the data processing methodology. The results found are then presented in chapter [5](#page-21-3) followed by an analysis and discussion in chapter [6.](#page-24-0) Chapter [7](#page-25-0) gives some final thoughts and summarises the study. Further research suggestions of subjects not covered in this study are presented in chapter [8.](#page-25-1) Finally, the report ends with a list of references used when writing this report.

# <span id="page-6-3"></span>**2 Methodology**

The paragraphs in the methodology section describe the type of study that is performed in this report. The approach to the study is also described.

# <span id="page-6-4"></span>**2.1 Choice of methodology**

There are traditionally two methods used when investigating problems similar to problems in this report. They are the qualitative and quantitative studies. When using a qualitative study a limited number of units are investigated to gain a deeper understanding of these units. A qualitative study most likely leads to a situation where the researcher's comprehension or interpretation of the information found serves as a basis for the results found in the study.

When using a quantitative study a vast number of units are studied with the purpose of gaining knowledge on a limited number of factors for each unit. Statistical analysis is most likely to take place of the data of interest such that different phenomenon can be explained using a selection of a certain population. A quantitative study can also be used to generalize and to represent other units in a similar population (Holme and Solvang, 1996).

The presented problems and the data used to find the results in this study use a quantitative methodology. It is possible to generalize the studied problems to similar populations (stock indices) and the results are based on statistical analysis of the collected data. A quantitative methodology to the study thus makes sense.

### <span id="page-7-0"></span>**2.2 Approach to the study**

In general there are two main approaches to data in a study. These approaches are the deductive approach and the inductive approach. When using a deductive approach the conclusions made for specific events are based on general principles and existing theories. Based on these known theories and principles hypothesis are derived and then further empirically tested for the cases studied. The choice and interpretation of data used are also influenced by the known theories and principles.

When using an inductive approach specific events are not derived from hypothesis derived from existing theories. The events are here studied without prior knowledge or influence of theories. A new theory is compiled and formulated using the results of the study (Davidsson, Patel, 1994).

The study presented in this report uses a deductive approach. Academic research in the field has been performed during several years investigating similar problems. Practices based on these investigations have been prepared and made accessible for the general public. A deductive approach in this case thus makes sense.

# <span id="page-7-1"></span>**3 Theoretical framework**

The paragraphs in the theoretical framework describe the theories used to form the results in this study.

### <span id="page-7-2"></span>**3.1 Asset returns, volatility and standard deviation**

The return of an asset today is defined as the difference between the closing price of the asset today and its closing price yesterday divided by yesterday's closing price.

$$
R_{t+1} = \frac{S_{t+1} - S_t}{S_t}
$$
 (1)

The log return is used widely as the output of the calculation is unit free. Log returns are thus suitable for comparison with other unit free returns. The return or daily geometric return, also called "log" return can be defined as the change in the logarithm of daily closing prices of an asset.

<span id="page-7-3"></span>
$$
R_{t+1} = \ln(S_{t+1}) - \ln(S_t)
$$
 (2)

The volatility is a measure of fluctuations in asset returns. An asset has a high volatility when the return fluctuates over a wide rage. When the return fluctuates over a small range the asset has a low volatility. Volatility can be seen as a risk measure or an uncertainty of asset return movements faced by participants in financial markets. The volatility is measured by the volatility or the standard deviation and is a measure of the asset returns dispersion over a specified time period.

<span id="page-8-1"></span>
$$
\sigma^2 = \frac{1}{N-1} \sum_{t=1}^{N} (R_t - \bar{R})^2
$$
 (3)

The standard deviation is simply the square root of the volatility.

$$
\sigma = \sqrt{\frac{1}{N-1} \sum_{t=1}^{N} (R_t - \bar{R})^2}
$$
 (4)

where:



#### <span id="page-8-0"></span>**3.2 Stylized facts of volatility and asset returns**

In order to model future volatility or the volatility of assets forming a portfolio or index it is important to consider the stylized facts. These facts have been found through vast academic research and seem to fit for most markets, time periods and types of assets.

1. It can be found that daily returns have very little autocorrelation. Thus, returns are almost impossible to predict by only looking at the past returns.

$$
Corr (R_{t+1}, R_{t+1-\tau}) \approx 0 \text{ for } \tau = 1, 2, 3, ..., 100 \tag{5}
$$

The correlation of returns with returns lagged from 1 to 100 days of the OMXS30 index and the MIB30 index are shown in [Figure 1](#page-9-0) and [Figure 2.](#page-9-1) When investigating these graphs it can be argued that the conditional mean fluctuates around zero and is fairly constant.



<span id="page-9-0"></span>The unconditional distribution of daily returns usually deviates from normality. This phenomenon is called excess kurtosis. [Figure 3](#page-9-2) and [Figure 4](#page-9-3) show a histogram along with the normal distribution. A closer look tells us that the histograms in these figures have fatter tails and have higher peaks around zero. The tails are usually fatter on the left side of the distribution. Fatter tails to the left of the distribution indicates higher probabilities of large losses compared to losses under the normal distribution. The QQ-plots in [Figure 5](#page-9-4) and [Figure](#page-9-5)  [6](#page-9-5) further shows the deviation from normality in the tails.



<span id="page-9-2"></span>Figure 3 – Daily OMXS30 returns against the normal distribution



<span id="page-9-4"></span>Figure 5 – QQ-Plots of daily OMXS30 returns against the normal distribution

<span id="page-9-1"></span>

<span id="page-9-3"></span>Figure 4 – Daily MIB30 returns against the normal distribution



<span id="page-9-5"></span>Figure 6 – QQ-plot of daily MIB30 returns against the normal distribution

- 2. Stock market movements tend to have more large drops in value than corresponding large increases. The market is thus negatively skewed. Once again this can be seen in [Figure 3](#page-9-2) and [Figure 4.](#page-9-3)
- 3. It is hard to statistically reject a zero mean of return. For short horizons such as daily the standard deviation of returns is given more importance. The OMXS30 index has a daily standard deviation of 1.1644% and a daily mean of 0.0281%. For the MIB30 index the numbers are 0.8804% and 0.0103% respectively.
- 4. It can be found that volatility measured by squared returns have a positive correlation with past returns, especially for short horizons such as daily or weekly. [Figure 7](#page-10-0) and [Figure 8](#page-10-1) shows the autocorrelation in squared returns for the OMXS30 index and the MIB30 index.



<span id="page-10-1"></span><span id="page-10-0"></span>5. The correlation tends to increase between assets. Especially in bearish markets or during market crashes. This phenomenon is also referred to as volatility clustering where extreme returns are followed by other extreme returns, although not necessarily of the same sign. [Figure 9](#page-10-2) an[d Figure 10](#page-10-3) shows the returns for the OMXS30 index and the MIB30 index over time where volatility clusters are present.

<span id="page-10-3"></span><span id="page-10-2"></span>

6. Equity and equity indices tend to have a negative correlation between their volatility and returns. This phenomenon is called the leverage effect and is due to the fact that a drop in stock price will increase the leverage of a firm under a constant capital structure, i.e. if debt stays constant.

These stylized facts are the most common facts mentioned in academic literatures. Other facts are also present. Regardless, it is important to know how these effects influence the financial data when using models for volatility forecasting. When considering these facts the daily return can be specified (Christoffersen, 2003).

<span id="page-11-4"></span><span id="page-11-2"></span>
$$
R_{t+1} = \mu_{t+1} + \varepsilon_{t+1} \tag{6}
$$

$$
\varepsilon_{t+1} = \sigma_{t+1} z_{t+1} \text{ and } z_{t+1} \sim \text{i.i.d D}(0,1) \tag{7}
$$

The random variable  $z_{t+1}$  is an innovation term which is assumed to be identically and independently distributed according to the distribution D(0,1). This distribution in this study is assumed to be the normal distribution with a mean equal to zero and volatility equal to one. The conditional mean of the return  $\mu_{t+1}$  is assumed to be zero (Christoffersen, 2003).

Researchers sometimes include an autoregressive term of the first order AR(1) or a Moving Average term of the first order MA(1) in the mean equation. A combination of these terms of the first order will be an ARMA(1,1) term. Higher orders of these terms AR(p), MA(q), are also possible to include in the mean equation.

<span id="page-11-3"></span>
$$
R_{t+1} = \mu_{t+1} + \sum_{i=0}^{p} \beta_{i,AR} AR(i) + \sum_{j=0}^{q} \beta_{j,MA} MA(j) + \varepsilon_{t+1}
$$
 (8)

By including these terms the model (mean equation) might capture dynamic features of the data better. The AR term states that a time-series current value depends on its current and previous value of a white noise error term. A time-series linear dependency on its own previous value is captured by the MA term (Brooks, 2008). The terms are selected using the Akaike Information Criterion (AIC), see paragraph [3.5.1.](#page-17-0)

#### <span id="page-11-0"></span>**3.3 Volatility modelling**

This study focus on Generalized Autoregressive Conditional Heteroskedastic (GARCH) models that are exploiting different aspects of the stylized facts described in paragraph 3.2 above. The following paragraphs will describe the different GARCH models in more detail. Other more simple models are also described as these are the foundation of the more complex GARCH models.

#### <span id="page-11-1"></span>**3.3.1 The Random Walk (RW) model**

The Random Walk (RW) model only considers the actual volatility in the present time period as a forecast for the volatility in the following time period. Thus, volatility is almost impossible to forecast based on historical information as the movement on the stock markets are seen as random.

$$
\sigma_{t+1}^2 = \sigma_t^2 \tag{9}
$$

12

where:

 $\sigma_{t+1}^2$ is the forecasted unconditional volatility based on the actual volatility at time t  $\sigma_t^2$ is the unconditional volatility at time t from equation [\(3\)](#page-8-1) with a 25 day sliding window, which is the average number of trading days in a month

#### <span id="page-12-0"></span>**3.3.2 The Moving Average (MA) model**

The Moving Average (MA) model states that if a period in time has a high volatility there is also a high probability of high volatility in the following time periods. Tomorrow's volatility is thus the simple average of the most recent observations.

<span id="page-12-3"></span>
$$
\sigma_{t+1}^2 = \frac{1}{m} \sum_{\tau=1}^m \sigma_{t+1-\tau}^2
$$
 (10)

where:



The drawback of this model is that extreme returns will influence the volatility for 1/m times the squared return for m time periods as equal weights are put on all past observations. Another drawback is that it can be difficult to select the number of observations.

### <span id="page-12-1"></span>**3.3.3 The Exponentially Weighted Moving Average (EWMA) model**

The Exponential Weighted Moving Average model (EWMA) or the exponential smoother, developed by JP Morgan considers the drawback with equal weights on past observations. Weights on the observations are declining exponentially over the time period. The influence on extreme returns is smoothed out compared to the Moving Average model.

$$
\sigma_{t+1}^2 = \lambda \sigma_t^2 + (1 - \lambda) \sigma_t^2 \tag{11}
$$

where:

 $\sigma_{t+1}^2$ is the forecasted conditional volatility based on information given at time t  $\lambda \sigma_t^2$ is the volatility at time t  $(1 - \lambda)\sigma_t^2$ is the is the forecasted unconditional volatility at time t from equation [\(10\)](#page-12-3)  $\lambda$  is the smoothing parameter

The smoothing parameter  $\lambda$  shall be estimated with values ranging from zero to one. When the exponential smoother model was developed it was found that a  $\lambda$ =0,94 was best for most assets for daily data (Jorion, 2001). This value has since then become an industry standard and is also used in this study.

### <span id="page-12-2"></span>**3.3.4 The ARCH(q) model**

R F Engle (1982) proposed a model where the conditional volatility was modelled with an AutoRegressive Conditional Heteroscedasticity (ARCH) processes. The term autoregressive here states that past events influence future events but with diminishing effect as time pass by. Engle used the fact that financial time series usually incorporates clusters of high and low volatility over longer time periods. The time series thus show heteroscedasticity. As the volatility usually is gathered in clusters Engle found no independent variable that explained this phenomenon. He therefore used a regression based model with past squared returns as a describing factor for volatility forecasting. The lag parameter for past squared returns is described with q. The most simplest ARCH(q) process is a ARCH(1) process where the squared return is lagged one time period.

$$
\sigma_{t+1}^2 = \omega + \sum_{i=1}^q \alpha_i \varepsilon_{t+1-i}^2
$$
 (12)

with  $\omega > 0$  and  $\alpha_i \geq 0$  for all  $i = 1$  to  $q$ .

where:



#### <span id="page-13-0"></span>**3.3.5 The GARCH(p,q) model**

The ARCH model was extended to a Generalized AutoRegressive Conditional Heteroscedasticity (GARCH) model by Tim Bollerslev (1986). This model not only considers past squared returns but also past volatilities to forecast the volatility. Thus the ARCH model was extended with an autoregressive term q for the volatility. The most simplest GARCH(p,q) process is a GARCH(1,1) process.

$$
\sigma_{t+1}^2 = \omega + \sum_{i=1}^p \alpha_i \varepsilon_{t+1-i}^2 + \sum_{j=1}^q \beta_j \sigma_{t+1-j}^2
$$
\n(13)

with  $\omega > 0$ ,  $\alpha_i \ge 0$  for all  $i = 1$  to  $p$ ,  $\beta_j \ge 0$  for all  $j = 1$  to  $q$  and (sum of  $\alpha_i$  + sum of  $\beta_j$ ) < 1.

where:



#### <span id="page-13-1"></span>**3.3.6 The EGARCH(p,q) model**

The GARCH(p,q) model does not consider the fact that a drop in asset return has a larger effect on the volatility then a corresponding rise in asset return. This effect is called the

leverage effect and is also argued in stylized fact number six above. The GARCH(p,q) model can be modified to deal with this phenomenon. The Exponential GARCH (EGARCH) model proposed by Nelson (1991) considers the asset returns effect on volatility differently depending on whether the return is positive or negative. An advantage with the model is that the volatility is always positive. The drawback with the model is that the forecasted volatility for lags larger than one cannot be calculated analytically. The most simplest EGARCH( $p,q$ ) process is a EGARCH $(1,1)$  process.

$$
\log \sigma_{t+1}^2 = \omega + \sum_{i=1}^p \alpha_i \left| \frac{\varepsilon_{t+1-i}}{\sigma_{t+1-i}} \right| + \sum_{k=1}^r \gamma_k \left| \frac{\varepsilon_{t-k}}{\sigma_{t-k}} \right| + \sum_{j=1}^q \beta_j \log(\sigma_{t+1-j}^2)
$$
(14)

with no restrictions on the parameters since the logarithm prevents  $\sigma_{t+1}^2$  to be negative.

where:



#### <span id="page-14-0"></span>**3.3.7 The GJR-GARCH(p,q) model**

Glosten, Jagannathan and Runkle (1993) proposed another extension to the GARCH model, which deals with the leverage effect. An indicator variable was here introduced to form the GJR-GARCH model. The indicator variable takes the value 1 if the return at time period t is positive. Otherwise the indicator variable takes the value 0. The most simplest GJR-GARCH (p,q) process is a GJR-GARCH(1,1) process.

$$
\sigma_{t+1}^2 = \omega + \sum_{i=1}^p \alpha_i \, \varepsilon_{t+1-i}^2 + \sum_{k=1}^r \gamma_k \varepsilon_{t+1-k}^2 \, I_{t+1-k}^- + \sum_{j=1}^q \beta_j \, \sigma_{t+1-j}^2 \tag{15}
$$

with  $\omega$  > 0,  $\alpha_i$  ≥ 0 for all  $i$  = 1 to  $p$ ,  $\beta_j$  ≥ 0 for all  $j$  = 1 to  $q$  , ( $\alpha_i+\gamma_k$ ) ≥ 0 for all  $i$  = 1 to  $p$  and  $k$ = 1 to  $r$  and (sum of  $\alpha_i$  + sum of  $\beta_j$  + 0.5 sum of  $\gamma_k$ ) < 1 for all  $i$  = 1 to  $p$ ,  $j$  = 1 to  $q$  and  $k$  = 1 to  $r$ .

where:

$$
\sigma_{t+1}^2
$$
 is the forecasted conditional volatility based on information given at time t  $\sigma_{t+1-j}^2$  is the volatility at time t+1-j



#### <span id="page-15-0"></span>**3.3.8 The APGARCH(p,q) model**

The Asymmetric Power GARCH (APGARCH) model was introduced by Ding, Granger and Engle (1993). It is similar to the GJR-GARCH model as the leverage effect is considered. The difference with the APGARCH models is that it allows the power к to be estimated. The power к usually takes the value 2 (Ding, Granger and Engle (1993)). The most simplest APGARCH(p,q) process is a APGARCH(1,1) process.

$$
\sigma_{t+1}^{\delta} = \omega + \sum_{i=1}^{p} \alpha_i (|\varepsilon_{t+1-i}| - \gamma_i \varepsilon_{t+1-i})^{\delta} + \sum_{j=1}^{q} \beta_j \sigma_{t+1-j}^{\delta}
$$
(16)

with δ > 0,  $|\gamma_i|$  ≤ 1 for all  $i$  = 1 to  $r$ ,  $\gamma_i$  = 0 for all  $i$  >  $r$  and  $r$  ≤  $p$ .

where:



The APGARCH model also includes other GARCH models as special cases. If restrictions are set on its parameters the below models can be specified:

- ARCH when  $\delta = 2$ ,  $\gamma_i = 0$  for all  $i = 1$  to  $p$  and  $\beta_i = 0$  for all  $j = 1$  to  $q$
- GARCH when  $\delta = 2$ ,  $\gamma_i = 0$  for all  $i = 1$  to  $p$
- GJR-GARCH when  $\delta = 2$

#### <span id="page-15-1"></span>**3.3.9 Other GARCH models**

There exist a variety of other GARCH models. The Threshold GARCH (TGARCH) model introduced by Zakoian (1994) and the Quadratic GARCH (QGARCH) model introduced by Sentana (1995) are similar to the GJR-GARCH model. They both consider the leverage effect. Other models as the Integrated GARCH (IGARCH) or the GARCH in mean (GARCH-M) are models that uses restrictions on the basic GARCH model or uses GARCH terms in the mean equation for the return. More complex GARCH models are the Regime Switching GARCH (RS-GARCH) model introduced by Gray (1996) or the fractionally integrated GARCH model (FIGARCH) introduced by Baillie, Bollerslev and Mikkelsen (1996), which forecasts volatility at longer time periods than daily or weekly. New GARCH models are proposed continuously but they all origin from the simple ARCH model proposed by Engel (1982). A list of many proposed models is found in the Glossary to ARCH (GARCH) also called as "alphabet-soup" (Bollerslev, 2008). The GARCH models mentioned in this paragraph will not be dealt with in this study.

### <span id="page-16-0"></span>**3.4 Maximum-Likelihood parameter estimation**

The GARCH models described in previous sections contain unknown parameters. These parameters must be estimated in order to fit the GARCH models properly to the historical (in-sample) data used for volatility forecasting. The method used to find the unknown parameters is based on Maximum Likelihood Estimation (MLE).

If equation [\(6\)](#page-11-2) is considered and the assumption of i.i.d. under the normal distribution, the likelihood or probability that  $R_t$  will occur is defined as  $l_t$  (Christoffersen, 2003).

$$
l_t = \frac{1}{\sqrt{2\pi\sigma_t^2}} \exp\left(-\frac{R_t^2}{2\sigma_t^2}\right) \tag{17}
$$

The joint likelihood of all squared returns within the time period is defined as  $L$ .

$$
L = \prod_{t=1}^{T} l_t = \prod_{t=1}^{T} \frac{1}{\sqrt{2\pi\sigma_t^2}} exp\left(-\frac{R_t^2}{2\sigma_t^2}\right)
$$
(18)

The unknown parameters are then selected to fit the data by maximizing the joint log likelihood of all squared returns within  $L$ .

$$
\max \log L = \max \sum_{t=1}^{T} \ln(l_t) = \max \sum_{t=1}^{T} \left[ -\frac{1}{2} \ln(2\pi) - \frac{1}{2} \ln(\sigma_t^2) - \frac{1}{2} \frac{R_t^2}{\sigma_t^2} \right] \tag{19}
$$

where:

 $L$  is the likelihood of all squared asset returns between time t and T  $l_t$ is the likelihood of the squared asset return at time t  $\sigma_t^2$ is the volatility at time t  $R_t^2$ is the squared asset return at time t

### <span id="page-16-1"></span>**3.5 Evaluation statistics**

Several statistical tests are performed in the software package EViews to determine which volatility models that are most suitable for the data used in the study. This paragraph

describes the Akaike Information Criterion (AIC) and the ARCH Lagrange multiplier (ARCH LM) test.

### <span id="page-17-0"></span>**3.5.1 Akaike Information Criterion (AIC)**

Models will fit a time-series as daily data from a stock index with different success depending on their specifications. Information criterion can be used as a guide when selecting the most appropriate model. The Akaike Information Criterion (AIC) consists of two parts. The first part is a test statistics derived from the maximum likelihood function. A penalty is then calculated in the second part for the loss of degrees of freedom from adding extra parameters (Brooks, 2008). The lower the information criteria the better the model fits the data.

$$
AIC = -\frac{2l}{T} + \frac{2k}{T}
$$
 (20)

where:



### <span id="page-17-1"></span>**3.5.2 ARCH Lagrange Multiplier (ARCH LM) test**

When using ARCH models for volatility forecasting it is appropriate to test if the time series consists of ARCH effects in the residuals. That is, if there is evidence of autoregressive conditional heteroskedasticity. If there is no evidence, ARCH models are not appropriate for volatility forecasting. The ARCH Lagrange multiplier (ARCH LM) test checks for ARCH effects in the residuals. When the ARCH models have been used and their parameters have been estimated for the in-sample period there should be as little evidence as possible left of any further ARCH effects. If there is still evidence of ARCH effects in the residuals there is more information in higher orders of the model that may explain the volatility better. A regression is used to compute two test statistics. The null hypothesis that there are no ARCH effects in the residuals up to order q is then tested (EViews, 2005).

$$
\varepsilon_t^2 = \beta_0 + \left(\sum_{i=1}^q \beta_i \,\varepsilon_{t-i}^2\right) + \nu_t \tag{21}
$$

where:

- $\varepsilon_t^2$ is the squared residual at time t
- $k$  is the number of estimated parameters
- $\beta_0$ is the intercept
- $\beta_i$ is the slope
- $v_t$ is the error term
- q is the non-negative order of the Moving Average ARCH term

The two test statistics computed for this test are the F-statistic and the OBS\*R-squared test statistic. From these test statistics p-values can be calculated, which is used to check if the null hypothesis can be rejected or not.

### <span id="page-18-0"></span>**3.6 Forecast volatility model evaluation**

When the volatility model has been selected and the parameters have been estimated it is necessary to check if the model performs as expected. This paragraph describes the evaluation process based on regression, the  $R^2$  test and the F-statistic test.

#### <span id="page-18-1"></span>**3.6.1 Out-of-sample check using regression**

A volatility model can be evaluated using regression. The squared return for the forecast period  $R_{t+1}^2$  is regressed on the volatility forecast from the volatility model itself.

$$
R_{t+1}^2 = b_0 + b_1 \sigma_{t+1}^2 + e_{t+1}
$$
 (22)

where:



The volatility forecast  $\sigma_{t+1}^2$  should be unbiased and efficient. It is unbiased if it has an intercept of  $b_0 = 0$  and efficient if the slope  $b_1 = 1$ . It shall be noted that as the squared return shows a high degree of noise the fit of the regression will be fairly low, typically around 5 - 10% (Christoffersen, 2003).

#### <span id="page-18-2"></span>**3.6.2 Intraday high and low prices**

Instead of using the squared return as a proxy for the volatility the range can be used. The range is the difference between the logarithmic high and low price of an asset during a trading day.

<span id="page-18-3"></span>
$$
D_t = ln(S_t^{High}) - ln(S_t^{Low})
$$
\n(23)

where:



Peter Christoffersen (2003) suggests that the proxy for the daily volatility shall have the following form.

<span id="page-18-4"></span>
$$
\sigma_{r,t}^2 = \frac{1}{4ln(2)} D_t^2 \approx 0.361 D_t^2 \tag{24}
$$

As in the out-of-sample check in the previous paragraph, regression is used for evaluating the forecast from the volatility model.

$$
\sigma_{r,t}^2 = b_0 + b_1 \sigma_{t+1}^2 + e_{t+1}
$$
\n(25)

where:



### <span id="page-19-0"></span>**3.6.3 Goodness of fit statistic, R<sup>2</sup> test**

When comparing forecasted models they have to be distinguished in terms of goodness of fit between the data and the model before they can be ranked. The  $R^2$  test is such a measure explaining how well a model containing explanatory variables actually explains variations in the dependent variable (Brooks, 2008). In other words to what extent the parameters and variables in the models used in this study explain the forecasted volatility on the OMXS30 and MIB30 stock indices. The R<sup>2</sup> test uses regression and will equal 1 if the fit is perfect and 0 if the fit is no better than the mean of the dependent variable (EViews, 2005).

#### <span id="page-19-1"></span>**3.6.4 The F-statistic from regression**

The F-statistic reported from a regression tests the hypothesis that all slope parameters  $(b_n)$ except the constant (c) and the intercept ( $b<sub>0</sub>$ ) is zero. That is, if the regressed parameters includes information from the dependent variable. The reported p-value Prob(F-statistic) is the marginal significance level of the F-test. If the p-value is significant the null hypothesis is rejected and the slope parameters  $(b_n)$  are significantly different from zero (EViews 2005).

$$
F = \frac{R^2/(k-1)}{(1 - R^2)/(T - k)}
$$
\n(26)

where:



### <span id="page-19-2"></span>**4 Data and data processing methodology**

The paragraphs in the data and data processing methodology section describe the data and the methods used in this study. Data can be divided in primary data and secondary data when used for research purposes. Primary data is data that researchers gather themselves through interviews or observations during the work of the study. Secondary data is data that have already been found by others and that can be extracted from databases, books or journals etc. (Holme and Solvang, 1996). This report only uses secondary data.

#### <span id="page-20-0"></span>**4.1 Description of the data**

The data consists of daily prices from the OMXS30 and MIB30 indices. Both OMXS30 and MIB30 list the 30 most valued companies on the Stockholm Stock Exchange and the Milano Stock Exchange. The time period for both OMXS30 and MIB30 indices is  $31<sup>st</sup>$  of October 2003 to 30<sup>th</sup> of June 2008. That is roughly five years giving 1171 sample points for OMXS30 and 1185 sample points for MIB30. Peter Christoffersen (2003) argues that a general rule of thumb is to use the last 1000 data points for volatility forecasting. The number of sample points differs between the indices due to different number of holidays and different dates when holidays occurs between Sweden and Italy. Included with the daily prices the data also contains the daily range, highest and lowest daily price. The data has been retrieved using the Bloomberg databases.

#### <span id="page-20-1"></span>**4.1.1 In- and out-of-sample periods**

The RW model only uses the previous sample point of the actual volatility to forecast the volatility. Thus the actual volatility for sample 1171 for the OMXS30 index and 1185 for the MIB30 index is used giving the first forecasted value in the out-of-sample period. The forecasting is then performed one step ahead throughout the out-of-sample period.

$$
\sigma_{t+1}^2 = \sigma_t^2 \quad \text{and } t_{\text{OMXS30}} = 1172, 1173, \dots, 1299 \tag{27}
$$

$$
\sigma_{t+1}^2 = \sigma_t^2 \quad \text{and } t_{\text{MIB30}} = 1186, 1187, \dots, 1312 \tag{28}
$$

For the MA model the moving average window is set to equal the number of sample points in the in-sample period. That is the moving average window for the OMXS30 index is 1171 and 1185 for the MIB30 index. The forecasting is then performed using these sample point sliding windows throughout the out-of-sample period.

$$
\sigma_{t+1}^2 = \frac{1}{1171} \sum_{\tau=1}^{1171} \sigma_{t+1-\tau}^2 \quad \text{and } t_{\text{OMXS30}} = 1172, 1173, \dots, 1299 \tag{29}
$$

$$
\sigma_{t+1}^2 = \frac{1}{1185} \sum_{\tau=1}^{1185} \sigma_{t+1-\tau}^2 \quad \text{and } t_{\text{MIB30}} = 1186, 1187, \dots, 1312 \tag{30}
$$

The EWMA model only requires 100 lags in order to include 99,8% of all the weights necessary to calculate tomorrow's volatility (Christoffersen, 2003). To forecast the first day of volatility in the out-of-sample period weights 1072 to 1171 are used for the OMXS30 index and weights 1086 to 1185 are used for the MIB30 index. The forecasting is then performed using the 100 sample point sliding windows throughout the out-of-sample period.

$$
\sigma_{t+1}^2 = \lambda \sigma_t^2 + (1 - \lambda) \sigma_t^2 \quad \text{and } t_{\text{OMXS30}} = 1172, 1173, \dots, 1299 \tag{31}
$$

$$
\sigma_{t+1}^2 = \lambda \sigma_t^2 + (1 - \lambda) \sigma_t^2 \quad \text{and} \quad t_{\text{MIB30}} = 1186, 1187, \dots, 1312 \tag{32}
$$

The in-sample period when the coefficients for the ARCH type models are estimated is equal to the time period of the retrieved data, 31 $^{\rm st}$  of October 2003 to 30 $^{\rm th}$  of June 2008. When the coefficients have been found and fitted to the data using MLE the coefficients are used to

forecast the volatility one day ahead during the period  $1^{\text{st}}$  of July 2008 to 30<sup>th</sup> of December 2008. This period is the out of sample period. The forecasting is thus made in a sliding window one day ahead throughout the out-of-sample period.  $1^{\text{st}}$  of July 2008 is the first day of forecasted volatility and 30<sup>th</sup> of December is the last day of forecasted volatility. Note that the out-of-sample period includes returns with high levels of volatility.

### <span id="page-21-0"></span>**4.2 Criticism of the sources**

The quality of the gathered data is very important in order to prevent biased results. It is therefore sound to treat the data with some scepticism. Typical questions to consider are the purpose of the data, when and where the data was gathered or even why the data exists. Further questions to consider are the circumstances when the data was collected, who the originator of the data is and its relation to the data (Davidson and Patel, 1994).

The data used for this study is gathered by Bloomberg and used by many institutions and agents in the financial markets. There is no reason to think that this data is not correct or biased in any way for purposes only known by Bloomberg.

# <span id="page-21-1"></span>**4.3 Validity and reliability**

When researching a topic it is important that the result reflects the questions asked about the topic. Validity thus ensures that this is fulfilled. Reliability deals with the consistency of a number of measurements and is determined based on the processing of the data. Reliability can be increased if the data is treated carefully to avoid errors. Validity implies reliability, but reliability does not imply validity (Holme and Solvang, 1996).

Bloomberg provides data to professional institutions and agents in financial markets all over the world. They have processes and rules that ensure correct data in their databases. If this were not the case they could not charge users for their services. The data is thus seen as reliable. Methods and calculations used in this study are presented in such a way that they could be replicated. A description and definition of what this study will find is also provided. The study can then be seen as valid.

# <span id="page-21-2"></span>**4.4 Data processing**

The data from the two indices are used to calculate other necessary data in order to further study the forecasting models. First, the prices are transformed to returns using equation [\(2\).](#page-7-3) The actual volatility is computed using equation [\(3\).](#page-8-1) Each of the MA, EWMA and RW models forecasted volatility are calculated using their respective formulas in paragraph [4.1.1.](#page-20-1) Finally the range and the range volatility are calculated using equation [\(23\)](#page-18-3) and [\(24\)](#page-18-4) respectively. All calculations are performed in Microsoft Excel. The data are then imported to Quantitative Micro Software EViews for further analysis. Note that the forecasting for the ARCH type models are performed in EViews and thus didn't have to be calculated in Excel.

# <span id="page-21-3"></span>**5 Results**

The data in the result paragraphs describe the results found on empirical data. First the mean equation is examined followed by the stylized facts and found model parameters. Then the in-sample results for the different models are presented followed by the forecast and the out-of-sample results of the model valuation.

### <span id="page-22-0"></span>**5.1 The mean equation**

The mean specification is found in equation [\(8\).](#page-11-3) Regression is used to find if an AR(p), MA(q) or ARMA(p,q) term shall be included in the mean equation used by the ARCH type models. The EViews Akaike Information Criterion (AIC) is used to find the best model for the mean equation for the in-sample period. The results are show i[n Table 4](#page-28-1) in [Appendix 1.](#page-28-0) For the OMXS30 index the mean equation shall follow an AR(1) process whereas an ARMA(5,5) process is more suitable for the MIB30 index.

### <span id="page-22-1"></span>**5.2 Statistics of the OMXS30 and MIB30 indices**

Evidences of the stylized facts are already discussed in paragraph 3.3. Most of these facts can also be measured statistically[. Table 1](#page-22-4) shows the result of the sample statistics.



Zero mean test – Tests the null hypothesis of a zero mean in the daily returns Lagrange Multiplier test – Tests if there are any ARCH effects in the daily return series ADF1 – Tests for unit root in the return series, ADF2 – Tests for unit root in the squared return series

<span id="page-22-4"></span>Table 1 – Sample statistics for the OMXS30 and MIB30 indices, Oct 31, 2003 – Jun 30, 2008

### <span id="page-22-2"></span>**5.3 Higher orders of GARCH**

To test for higher orders of GARCH models the AIC criterion from EViews is used once again to find the best models. The results are shown in [Table 5](#page-29-1) through [Table 8](#page-32-1) in [Appendix 2](#page-29-0) through [Appendix 5.](#page-32-0) For higher orders, GARCH(7,2), EGARCH(3,7), GJR-GARCH(4,8) and APGARCH(2,7) are the most suitable models for the OMXS30 index. For the MIB30 index, GARCH(8,8), EGARCH(4,6), GJR-GARCH(2,4) and APGARCH(3,6) are the most suitable models.

An ARCH LM test was performed in order to ensure that there are no or few ARCH effects left in the residuals. The results of the ARCH LM test are shown i[n Table 9](#page-33-1) in [Appendix 6.](#page-33-0) All models passed the test. Thus, the models have been successful in removing any remaining ARCH effects in the data for both indices.

### <span id="page-22-3"></span>**5.4 Estimated parameters**

[Table 10](#page-34-1) and [Table 11](#page-35-1) in [Appendix 7](#page-34-0) shows the estimated parameters for the first order GARCH and the order higher than one that best fitted the data.

### <span id="page-23-0"></span>**5.5 Forecast model evaluation**

The one day ahead forecast for each model was evaluated with equation [\(24\)](#page-18-4) as a proxy. An  $R^2$  statistic was used to rank the models. The regression parameters along with the p-value from the F-test are also presented. The results are shown in [Table 2](#page-23-1) and [Table 3.](#page-23-2)



The R<sup>2</sup> statistics are applied over the period 01/07/2008 to 30/12/2008. Model rankings are obtained by expressing the R<sup>2</sup> statistics as a ratio relative to the worst performing model. The best model is marked in bold.



<span id="page-23-1"></span>Table 2 – Model ranking from forecasting daily volatility for the OMXS30 index.

The R<sup>2</sup> statistics are applied over the period 01/07/2008 to 30/12/2008. Model rankings are obtained by expressing the R<sup>2</sup> statistics as a ratio relative to the worst performing model. The best model is marked in bold.

<span id="page-23-2"></span>Table 3 – Model ranking from forecasting daily volatility for the MIB30 index.

The results show that the EGARCH(1,1) model performed best during the forecast period for both indices. As the actual volatility is unobserved the range from equatio[n \(24\)](#page-18-4) is plotted as a proxy against the forecasted volatilities to graphically show each models performance. The graphs are shown in [Figure 11](#page-36-1) through [Figure 34](#page-39-0) in [Appendix 9](#page-36-0) and [Appendix 10.](#page-38-0)

# <span id="page-24-0"></span>**6 Analysis and discussion**

There is evidence in [Figure 1](#page-9-0) and [Figure 2](#page-9-1) of insignificant correlations for the indices returns. In [Figure 7](#page-10-0) and [Figure 8](#page-10-1) the autocorrelations in squared returns are significant and persistent. The values found from the Ljung-Box Q-statistics and Engle's Lagrange multiplier ARCH test further supports these findings. The Ljung-Box Q-statistics determines that the null hypothesis of no autocorrelation up to lag 10 cannot be rejected. There is also strong evidence of autoregressive conditional heteroskedasticity determined by the Lagrange multiplier ARCH test. Daily returns deviate from normality. This is shown in [Figure 3](#page-9-2) and [Figure 4](#page-9-3) and is also supported by the values of excess kurtosis and by the Jarque-Bera normality test. The Jarque-Bera null hypothesis is rejected at the 1% significance level. The indices are asymmetric or negatively skewed, which once again is supported by the values of excess kurtosis and by the values of skewness. The assumption of a zero mean cannot be rejected at the 1% significance level for both indices. The null hypothesis states that the return has a mean of zero. [Figure 9](#page-10-2) and [Figure 10](#page-10-3) also show evidence of volatility clustering. The augmented Dickey-Fuller test shows that there are no unit roots in return and squared return series for both indices. A time-series as daily data from a stock index is stationary if the mean and volatility do not depend on time. If a unit root exists it could lead to the nonexistence of a stationary version of the volatility and lead to infinite volatility. These facts suggest that the GARCH models are preferred for more simple models. The GARCH models would also profit from having the innovation term from equation [\(7\)](#page-11-4) to follow a distribution with fatter tails such as the student-t distribution. The models would then fit the data better. Although a better fit in the in-sample period would not necessarily mean a better forecast in the out-of-sample period. This can be seen from [Table 2](#page-23-1) an[d Table 4](#page-28-1) where more or better information, in this case higher orders, doesn't improve the forecasting.

The higher orders of GARCH models are not superior to the lower orders. In fact no higher order GARCH model outperforms the lower order models. This can also be seen in [Appendix](#page-33-0)  [6](#page-33-0) where all or most ARCH effects have been removed by the models as the p-values for the test values reject the null hypothesis of existence of remaining ARCH effects. Thus, the volatility has been correctly specified for all GARCH models. Many GARCH models of higher orders have insignificant parameters. This simply means that if these parameters were zero they would not affect the outcome of the volatility from the model. Unfortunately many restrictions are also violated, i.e.  $ω$ ,  $α$  or  $β$  parameters are negative. For the OMXS30 index the Maximum Likelihood Estimation does not converge on some occasions, which means that more information is necessary to add to the model specifications. A conclusion is therefore that these models are less attractive to use even though they outperform the more simple models. Interestingly the Random Walk model outperforms the MA, EWMA and ARCH(1) models for both indices. The RW model does however lack information from the range, which can be seen by the slope parameter  $b_1$  compared to the GARCH models. GARCH models including the leverage effect seem to perform better. In particular the EGARCH(1,1) models performs best for both indices closely followed by the GJR-GARCH(1,1)

and APGARCH(1,1) specifications. As shown in [Table 2](#page-23-1) and [Table 3](#page-23-2) they are also unbiased, fairly efficient and include information about the range. If there was no information about the range the models lack representation of the underlying data on which the volatility is based.

# <span id="page-25-0"></span>**7 Summary**

This study has examined various models capability to predict volatility on the OMXS30 and MIB30 indices. The models, which were tested included the Random Walk, the Moving Average, the Exponential Weighted Moving Average, ARCH, GARCH, EGARCH, GJR-GARCH and APGARCH. The results suggest that GARCH models give better volatility forecasts than more simple models. Although the simple Random Walk model outperforms both ARCH, MA and EWMA models. In particular GARCH models, which includes the leverage effect performs best for the period 31 $^{\text{st}}$  of October 2003 to 30 $^{\text{th}}$  of December 2008. Thus, the GARCH framework should be preferred to simpler volatility models when used for further risk management purposes.

# <span id="page-25-1"></span>**8 Further research suggestions**

Regression has been used to evaluate the volatility models in this study. Instead of a regression approach one could use various error statistics to see if results differ. These error statistics could be Mean Error (ME), Mean Absolute Error (MAE), Root Mean Square Error (RMSE) or a Mean Absolute Percentage Error (MAPE) just to mention a few.

The study has focused on the assumption of normally distributed innovation terms. As seen in paragraph [5.2](#page-22-1) this assumption does not reflect reality. It could therefore be interesting to evaluate the performance of the volatility models under the assumptions of a Student-t, Skewed Student-t, Gereralized Error Distribution or any other distribution of choice.

Only a one day forecast horizons has been investigated in this study. It could be interesting to see how the GARCH models perform under longer forecast horizons. More specifically the regime switching models or fractionally integrated models could be investigated as these handle longer forecast horizons better than the GARCH models presented in this study.

For risk management purposes a volatility forecast using a volatility model is not really useful unless the results from the forecast are used to calculate risk measures. In particular such risk measures are the Value at Risk (VaR) or the Expected Shortfall (ES).

For further readings in the field of volatility forecasting it's recommended to read the book Elements of Financial Risk Management by Christoffersen (2003). The Gloria Mundi (2009) website contains many interesting articles and research papers free to download. There is also a very interesting research paper written by Brailsford and Faff (1996), which evaluates many volatility forecasting techniques. Finally, the research papers mentioned in the references, especially from Engle and Bollerslev are strongly recommended for further readings.

### <span id="page-26-0"></span>**9 References**

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# <span id="page-28-0"></span>**Appendix 1 – AIC results for the mean equation**

<span id="page-28-1"></span>Table 4 – Akaike Information Criterion for the mean equation. The best fitted ARMA terms models are shown in bold.



# <span id="page-29-0"></span>**Appendix 2 – AIC results of the GARCH(p,q) model**

\* MLE convergence not achieved after 500 iterations

<span id="page-29-1"></span>Table 5 – AIC test for higher orders of GARCH for the OMXS30 and MIB30 indices. The lowest order and best fitted models are shown in bold.

<span id="page-30-0"></span>



\* MLE convergence not achieved after 500 iterations

Table 6 – AIC test for higher orders of EGARCH for the OMXS30 and MIB30 indices. The lowest order and best fitted models are shown in bold.

<span id="page-31-0"></span>



\* MLE convergence not achieved after 500 iterations

Table 7 – AIC test for higher orders of GJR-GARCH for the OMXS30 and MIB30 indices. The lowest order and best fitted models are shown in bold.

<span id="page-32-0"></span>



\* MLE convergence not achieved after 500 iterations

<span id="page-32-1"></span>Table 8 – AIC test for higher orders of APGARCH for the OMXS30 and MIB30 indices. The lowest order and best fitted models are shown in bold.



# <span id="page-33-0"></span>**Appendix 6 – ARCH LM test for the GARCH models**

<span id="page-33-1"></span>Table 9 – ARCH LM test to check for remaining ARCH effects.



# <span id="page-34-0"></span>**Appendix 7 – Estimated parameters for the best fitted GARCH models for the OMXS30 index**

(\*\*) indicates statistical significance at the 1% level, (\*) indicates statistical significance at the 5% level (x) indicates no statistical significance

<span id="page-34-1"></span>Table 10 – Estimated parameters for the first order GARCH and the order higher than one that best fitted the data for the OMXS30 index.



# <span id="page-35-0"></span>**Appendix 8 – Estimated parameters for the best fitted GARCH models for the MIB30 index**

(\*\*) indicates statistical significance at the 1% level, (\*) indicates statistical significance at the 5% level (x) indicates no statistical significance

<span id="page-35-1"></span>Table 11 – Estimated parameters for the first order GARCH and the order higher than one that best fitted the data for the MIB30 index.

# <span id="page-36-0"></span>**Appendix 9 – Forecasted volatility with the range as a proxy for the OMXS30 index**



<span id="page-36-1"></span>Figure 11 – Forecasted RW volatility for the OMXS30 index



Figure 13 – Forecasted EWMA volatility for the OMXS30 index





Figure 12 – Forecasted MA volatility for the OMXS30 index



Figure 14 – Forecasted ARCH(1) volatility for the OMXS30 index



for the OMXS30 index



Figure 17 – Forecasted EGARCH(1,1) volatility for the OMXS30 index









Figure 18 – Forecasted EGARCH(3,7) volatility for the OMXS30 index



Figure 20 – Forecasted GJR-GARCH(4,8) volatility for the OMXS30 index



# <span id="page-38-0"></span>**Appendix 10 – Forecasted volatility with the range as a proxy for the MIB30 index**



Figure 23 – Forecasted RW volatility for the MIB30 index





for the MIB30 index



Figure 24 – Forecasted MA volatility for the MIB30 index



the MIB30 index





Figure 29 – Forecasted EGARCH(1,1) volatility for the MIB30 index









Figure 30 – Forecasted EGARCH(4,6) volatility for the MIB30 index



Figure 32 – Forecasted GJR-GARCH(2,4) volatility for the MIB30 index

<span id="page-39-0"></span>